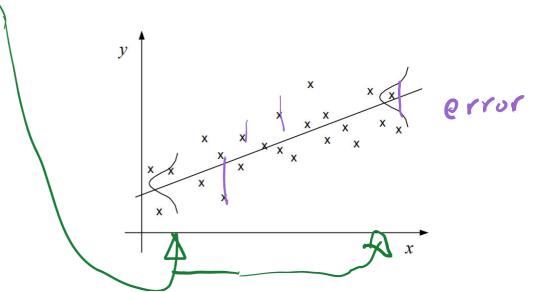


Today's lecture:

Linear regression (MLE + conditional Gaussians)

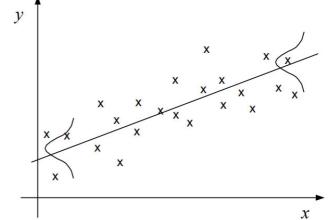
Regression

- Supervised learning: data pairs $D = \{(x_i, y_i)\}$, where, x_i may be discrete and/or continuous.
- Regression: label, y_i is a real-valued, e.g., $y_i \in \mathbb{R}$.
- Formally, want p(y|x) the conditional pdf.
- "Point" prediction is then $\hat{y} = E_Y[p(Y|X = x)]$.



Regression examples

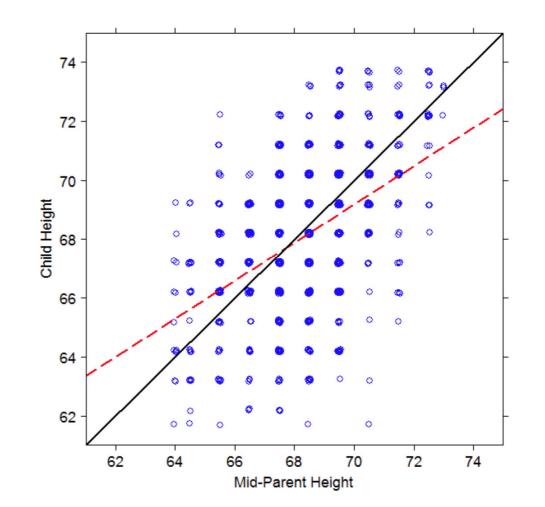
- Covid infection rates from zip code and vaccination rate, etc.
- How much a particular protein will bind to a drug target.
- A person's blood pressure from their genetics.
- Tracking object location in video at the next time-step.
- Housing prices, crime rates, stock prices, etc.
- Earliest regression: Legendre in 1805, and Gauss in 1809, both estimating orbits of bodies about the sun.



History of the term "Regression"

• Sir Francis Galton (1822-1911) "regression to the mean".

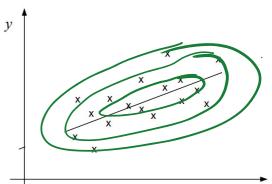
"It appeared from these experiments that the offspring did not tend to resemble their parents in size, but always to be more mediocre than they – to be smaller than the parents, if the parents were large; to be larger than the parents, if the parents were small."





Possible Regression Tactics (to estimate p(y|x))

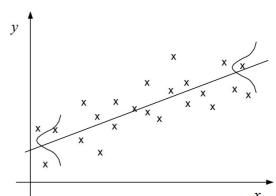
- Our data are drawn from some distribution, $(X, Y) \sim p(x, y)$.
- What are possible strategies to estimate p(y|x)?
- 1. Estimate $p(x, y|\theta)$ (e.g. MVG for RVs X, Y), and then use fitted model to compute $p(y|x, \hat{\theta}) = \frac{p(y, x|\hat{\theta})}{p(x|\hat{\theta})} = \frac{p(y, x|\hat{\theta})}{\int_{y} p(y, x|\hat{\theta}) dy}$.



Possible Regression Tactics (to estimate p(y|x))

- Our data are drawn from some distribution, $(X_i, Y_i) \sim p(x, y)$.
- What are possible strategies to estimate p(y|x)?
- 1. Estimate $p(x, y|\theta)$ (e.g. MVG for RVs X, Y), and then use fitted model to compute $p(y|x, \hat{\theta}) = \frac{p(y, x|\hat{\theta})}{p(x|\hat{\theta})} = \frac{p(y, x|\hat{\theta})}{\int_{y} p(y, x|\hat{\theta}) dy}$.
- 2. Consider the inputs to be fixed, and model only the output as a RV. That is, directly model the conditional $p(y|x,\hat{\theta})$.

"generative", vs. *"discriminative"*



Linear Regression

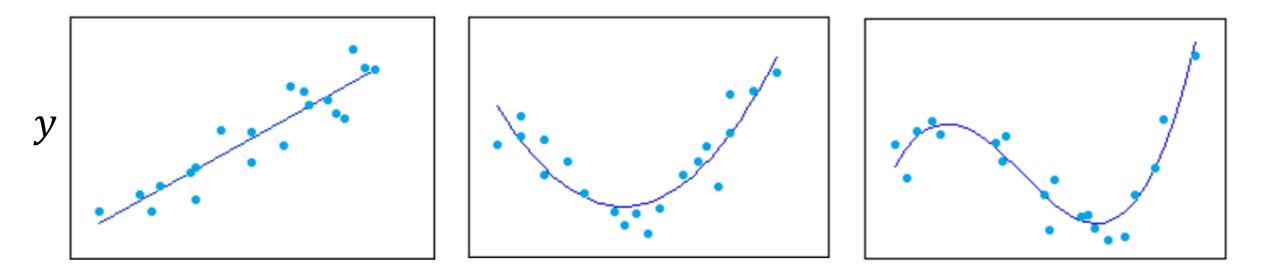
- Takes the discriminative approach.
- Predictions are a linear function of the <u>parameters</u>:

 $\hat{y} = E_Y[p(y|x)] = w^T x + w_0$, for $w, x \in \mathbb{R}^d$.

- w_0 is called the "offset"/"bias"/"intercept".
- Instead of a bias, we can make an extra feature that is always 1.
- Now use x' = [x, 1] and $\hat{y} = w^T x'$.

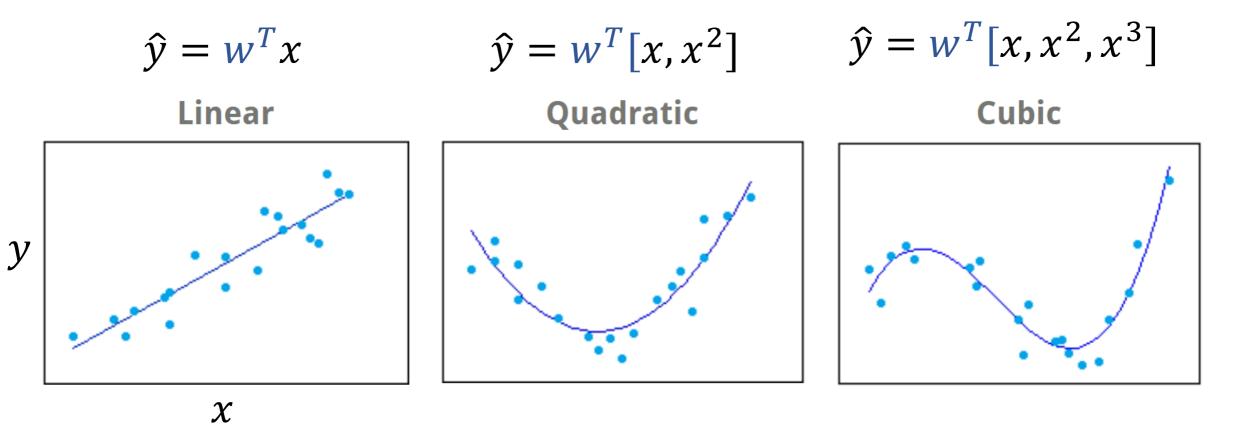
How useful can a linear model be?!

Which of these curves can be modelled by linear regression? $\hat{y} = E_Y[p(y|x)] = w^T x + w_0$, for $w, x \in \mathbb{R}^d$



https://statisticsbyjim.com/regression/curve-fitting-linear-nonlinear-regression/

How useful can a linear model be?! $w, x \in \mathbb{R}^1$



For full generality, $x \in \mathbb{R}^D$ need the cross-terms and bias terms, e.g., quadratic $[1, x_1, x_2, x_1^2, x_2^2, x_1x_2]$.

https://statisticsbyjim.com/regression/curve-fitting-linear-nonlinear-regression/

Basis expansion of raw input space

 $x \in \mathbb{R}^{d=2} = [x_1, x_2] \rightarrow [1, x_1, x_2, x_1x_2, x_1^2, x_2^2] \in \mathbb{R}^{k=6}$ Polynomial expansion of order 2 (i.e. quadratic)

• Denote *basis expansion* of the (raw) input features: $\Phi(x): \mathbb{R}^d \to \mathbb{R}^k$.

For d = 1, polynomial expansions of order k = 2, 3:

- For a quadratic expansion, $\Phi(x) = [1, x, x^2]$, and k = 2.
- For a cubic expansion, $\Phi(x) = [1, x, x^2, x^3]$, and k = 3.
- For a linear/identity basis expansion, $\Phi(x) = x$, and k = d.

Basis expansion of raw input space

Basis functions are pre-determined, so just a notational change: $\hat{y} = E_Y[p(y|x)] = w^T \Phi(x)$, for $w \in \mathbb{R}^k$, $x \in \mathbb{R}^d$

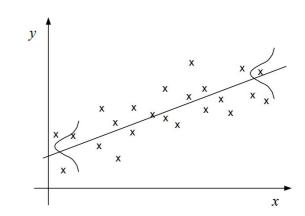
In this lecture, for simplicity of notation, we will assume that this expansion has already been done, and just write $\hat{y} = w^T x$.

Many basis possible functions!

Pdynomials:
$$f(x) = w_1 + w_2 x(1) + w_3 x(2) + w_3 x(1)^2 + w_3 x(2) + w_3 x(1) x(1) + w_3 x(1) x(2) + w_3 x(1) + w_$$

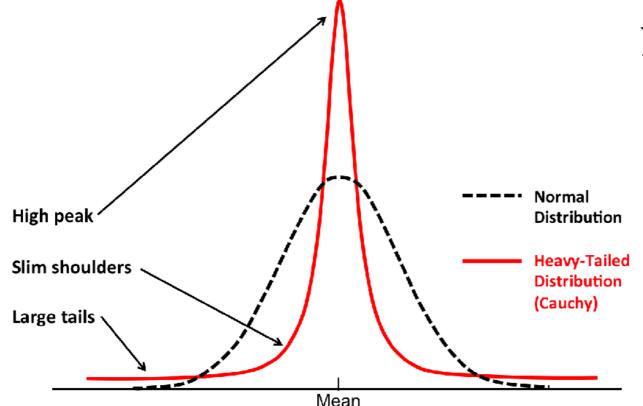
Specific form of linear regression

- So far we said linear regression is $\hat{y} = E_Y[p(y|x)] = E[p(y|x)] = w^T x$.
- But what do we use for p(y|x)?
- Standard linear regression uses a Gaussian $p(y|x) = N(y|w^T x, \sigma^2)$.
- Equivalent to $Y = w^T x + \epsilon$, with $\epsilon \sim N(0, \sigma^2)$.
- Which is equivalent to $Y w^T x = \epsilon \sim N(0, \sigma^2)$.
- Alternate forms give "heavier tails" to the distribution of the "residual".



Aside: heavy-tailed distribution

- More of the mass lies further from the center of mass.
- Heavy-tailed noise models outliers better than a Gaussian.



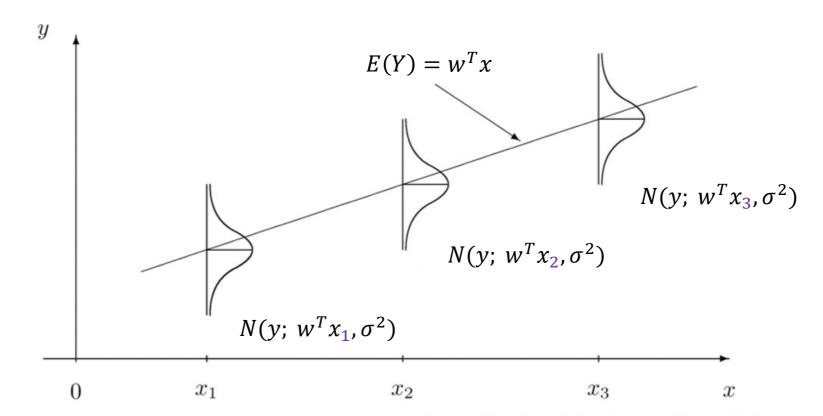
$$Y - w^{T}x = \epsilon \sim N(0, \sigma^{2})$$

$$Y - w^{T}x = \epsilon \sim Cauchy(0, \sigma^{2})$$

https://www.semanticscholar.org/paper/The-Normal-Distribution-.-.-or-Another-Brandenburger-Weston/b0a227e36d655c0dceaeb774225e0b4fb6cda8ed/figure/2

Gaussian linear regression, $p(y|x) = N(y|w^T x, \sigma^2)$

For every value X = x, the target variable, Y, takes on a Gaussian distribution with the same variance, σ^2 :



https://sdcastillo.github.io/PA-R-Study-Manual/introduction-to-modeling.html

How will we fit the regression model, $p(y|x) = N(y|w^Tx, \sigma^2)$? MLE: $\theta_{MLE} = (w_{MLE}, \sigma_{MLE}^2) = \arg \max_{(w,\sigma^2)} \log p(D = \{(x_i, y_i)_{i=1}^n\} | \theta)$ $= \arg \max_{(w,\sigma^2)} \sum_{i=1}^n \log p(y_i | x_i, \theta)$ $= \arg \max_{(w,\sigma^2)} \sum_{i=1}^n \log N(y_i | w^T x_i, \sigma^2)$ $= \arg \max_{(w,\sigma^{2})} \sum_{i=1}^{n} \log \left(\frac{1}{\sqrt{2\pi\sigma^{2}}} \right) - \frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (y_{i} - w^{T} x_{i})^{2}$ $= \arg \max_{(w,\sigma^{2})} n \log \left(\frac{1}{\sqrt{2\pi\sigma^{2}}}\right) - \frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (y_{i} - w^{T} x_{i})^{2}$

If we assume σ^2 is known...

If we assume σ^2 is known...

$$(w_{MLE}, \sigma_{MLE}^2) = \arg\max_{(w,\sigma^2)} \log\left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) - \frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - w^T x_i)^2$$

... then estimating w above is the same as

$$= \arg\min_{w} \sum_{i=1}^{n} (y_i - w^T x_i)^2$$
$$= \arg\min_{w} \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

"least squares" loss function!

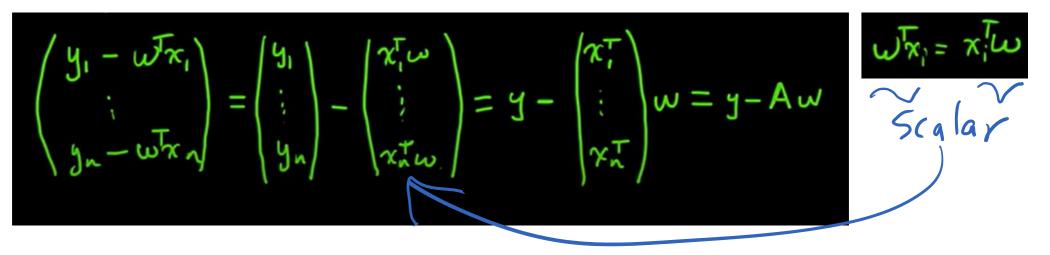
$$p(y|x) = N(y|w^T x, \sigma^2)$$

n

$$w_{\text{MLE}} = \arg\min_{w} \sum_{i=1}^{n} (y_i - w^T x_i)^2$$

Lets re-write this loss in "vectorized" form:

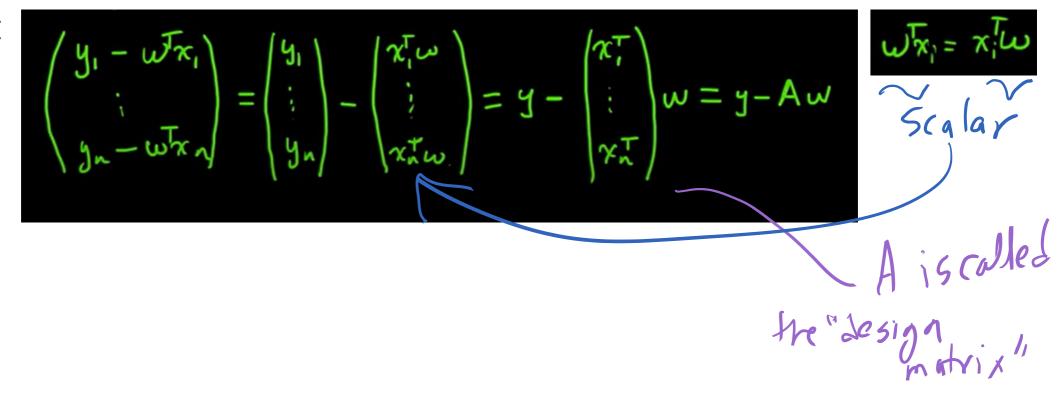
First, define:



$$w_{\text{MLE}} = \arg\min_{w} \sum_{i=1}^{n} (y_i - w^T x_i)^2$$

Lets re-write this loss in "vectorized" form:

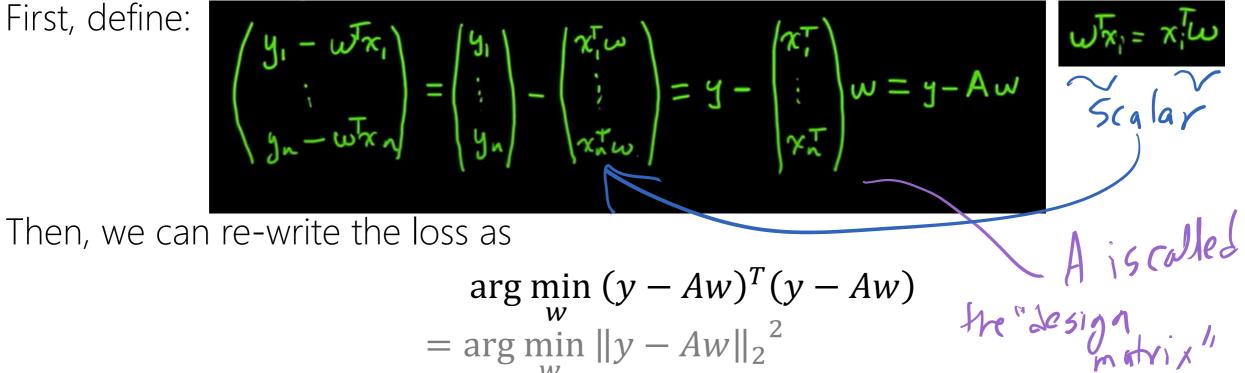
First, define:



$$\arg\min_{w} \sum_{i=1}^{n} (y_i - w^T x_i)^2$$

Lets re-write this loss in "vectorized" form:

First, define:



So want to minimize

$$\mathcal{L} = (y - Aw)^T (y - Aw) \quad (y \in \mathbb{R}^n, A \in \mathbb{R}^{n \times d}, w \in \mathbb{R}^{d \times 1}$$

= $(y^T - (Aw)^T)(y - Aw)$
= $y^T y - w^T A^T y - y^T Aw + w^T A^T Aw$
= $y^T y - 2w^T A^T y + w^T A^T Aw$

To minimize, we want to set $\frac{\partial \mathcal{L}}{\partial w} = 0$.

This is most easily achieved by using the rules of vector calculus, so lets do a quick refresher.

$$p(y|x) = N(y|w^T x, \sigma^2)$$

XT

Refresher on vector calculus

Some "rules" for taking gradients with respect to vectors.

• e.g., for vectors $a, b \in \mathbb{R}^{d \times 1}$, so that $a^T b \in \mathbb{R}$,

$$\frac{\partial(a^Tb)}{\partial a_j} = \frac{\partial(a_1b_1 + a_2b_2 + \cdots + a_db_d)}{\partial a_j} = b_j$$

Thus,

$$\frac{\partial (a^T b)}{\partial a} = \frac{\partial (a_1 b_1 + a_2 b_2 + \dots + a_d b_d)}{\partial a} = b \in \mathbb{R}^{d \times 1}$$
$$\frac{\partial (b^T a)}{\partial a}$$

(not true for ab^T which is a matrix, be careful!)

$$=\frac{\partial(b^Ta)}{\partial a}$$

Useful cheat sheet: https://cs.nyu.edu/~roweis/notes/matrixid.pdf

Refresher on vector calculus

For vector $x \in \mathbb{R}^{d \times 1}$, and matrix $\Sigma \in \mathbb{R}^{d \times d}$

$$\frac{\partial x^T \Sigma x}{\partial x} = (\Sigma + \Sigma^T) x$$

Thus, if Σ is symmetric such that $\Sigma = \Sigma^T$ then

$$\frac{\partial x^T \Sigma x}{\partial x} = 2\Sigma x$$

(similar to the scalar version: $\frac{\partial (ax^2)}{\partial x} = 2ax$)

So want to minimize

$$\mathcal{L} = (y - Aw)^T (y - Aw)$$

$$= (y^T - (Aw)^T)(y - Aw)$$

$$= y^T y - w^T A^T y - y^T Aw + w^T A^T Aw$$

$$= y^T y - 2w^T A^T y + w^T A^T Aw$$

To minimize, we want to set $\frac{\partial \mathcal{L}}{\partial w} = 0$.

 $(y \in \mathbb{R}^n, A \in \mathbb{R}^{n \times d}, w \in \mathbb{R}^{d \times 1}) \qquad \begin{pmatrix} \chi, \\ \vdots \\ \chi, \\ \chi \end{pmatrix} = A$

So want to minimize

$$\mathcal{L} = (y - Aw)^{T}(y - Aw)$$

$$= (y^{T} - (Aw)^{T})(y - Aw)$$

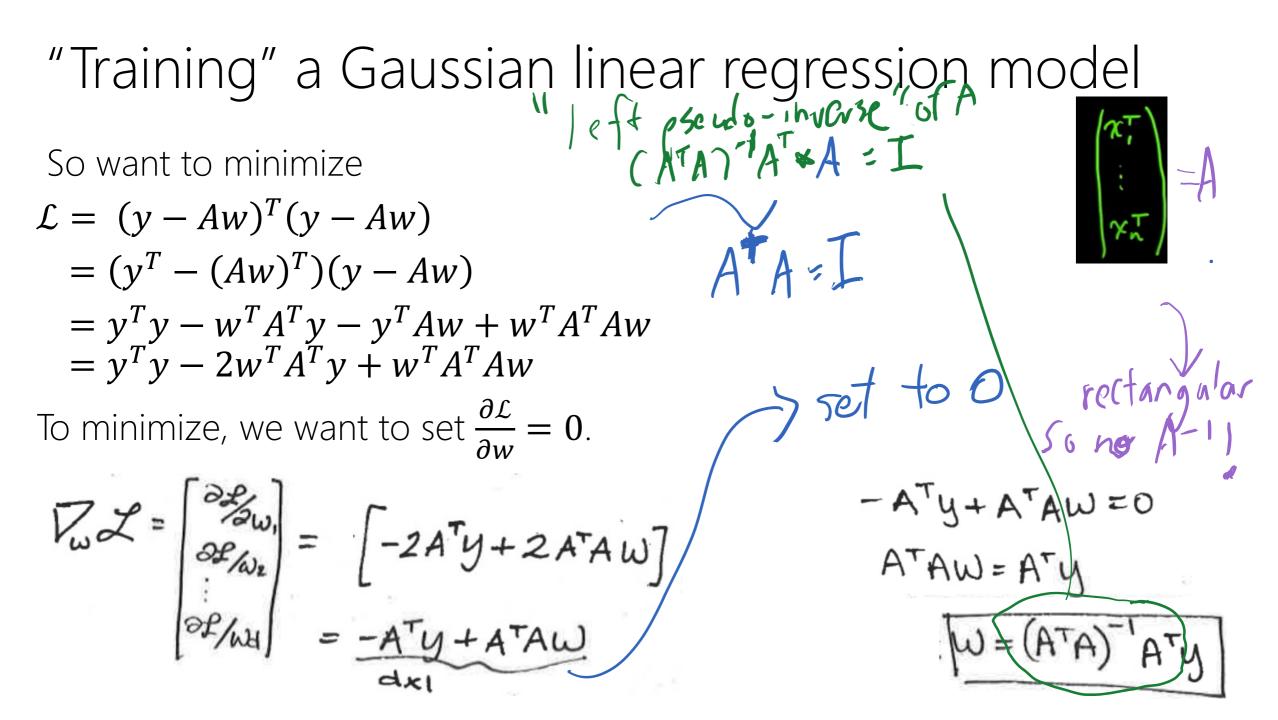
$$= y^{T}y - w^{T}A^{T}y - y^{T}Aw + w^{T}A^{T}Aw$$

$$= y^{T}y - 2w^{T}A^{T}y + w^{T}A^{T}Aw$$

To minimize, we want to set $\frac{\partial \mathcal{L}}{\partial w} = 0$.

$$V_{\omega} \mathcal{L} = \begin{bmatrix} \partial \mathcal{L}_{\omega} \\ \partial \mathcal{L}_{\omega} \\ \partial \mathcal{L}_{\omega} \end{bmatrix} = \begin{bmatrix} -2A^{T}y + 2A^{T}Aw \end{bmatrix}$$
$$= -A^{T}y + A^{T}Aw$$
$$= -A^{T}y + A^{T}Aw$$
$$= -A^{T}y + A^{T}Aw$$

$$\begin{pmatrix} \mathbf{x}^{\mathsf{T}} \\ \vdots \\ \mathbf{x}^{\mathsf{T}} \\ \mathbf{x}^{\mathsf{T}} \end{pmatrix} = A$$



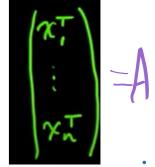
- We still need to check if the critical point, $w = (A^T A)^{-1} A^T y$ is minimum of the squared error loss.
- Recall $\nabla_w \mathcal{L} = -2A^T y + 2A^T A w$
- So Hessian matrix $(\nabla_w^2 \mathcal{L})$ is $2A^T A$. When is $A^T A$ PD?
- When the features are independent (when it has full rank).
- σ^2 from MLE as well is just the mean squared residual, $\sigma^2 = \frac{1}{N} \sum_i (y_i - w^T x)^2$.

$$\nabla_{\omega} \mathcal{L} = \begin{bmatrix} \partial_{\omega} \partial_{\omega} \\ \partial_{\omega} \partial_{\omega} \\ \partial_{\omega} \partial_{\omega} \end{bmatrix} = \begin{bmatrix} -2A^{T}y + 2A^{T}AW \end{bmatrix} \\
 \frac{\partial_{\omega} \partial_{\omega}}{\partial_{\omega} \partial_{\omega}} = -A^{T}y + A^{T}AW \\
 \frac{\partial_{\omega} \partial_{\omega}}{\partial_{\omega}$$

$$-A^{T}y + A^{T}AW = 0$$

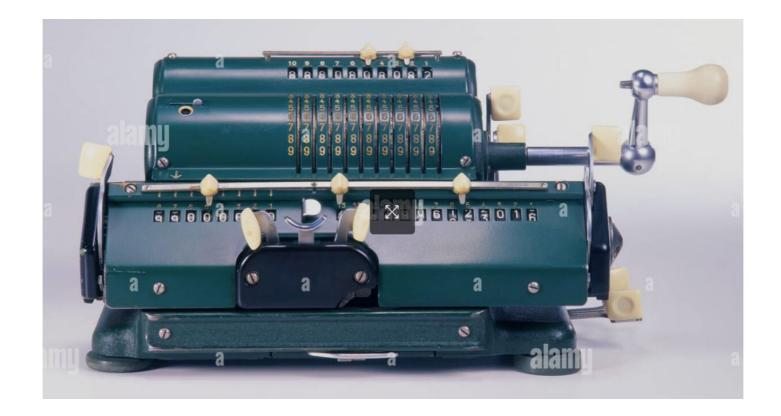
$$A^{T}AW = A^{T}y$$

$$W = (A^{T}A)^{T}A^{T}y$$



Regression in 1950s

Electromechanical desk "calculators" were used, and it could take up to 24 hours to receive the result from one regression.

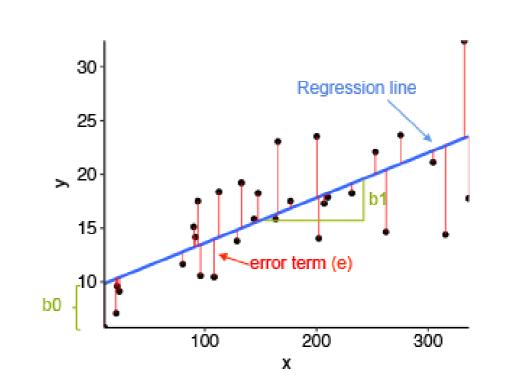


Geometric view of linear regression

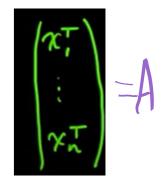
- We're trying to predict all our training data labels correctly, such that $y_i = w^T x_i$ for all $i \in [1 \dots n]$.
- In vector form, this means we're looking for

$$\begin{pmatrix} y_{i} \\ y_{2} \\ \vdots \\ y_{N} \end{pmatrix} = A_{W} = \begin{bmatrix} x_{i} \\ x_{2} \\ \vdots \\ y_{N} \end{bmatrix} W$$

Generally not possible because of noise; or incorrect model (e.g. all features).



https://www.analyticsvidhya.com/blog/2021/06/25-questions-to-test-your-skills-on-linear-regression-



Geometric view of linear regression, y = Aw

- So lets think about the error vector ($\mathbf{e} \equiv y \hat{y} = y Aw$, $\in \mathbb{R}^{n \times 1}$).
- A "good" setting of w minimizes the length of e.
- The length is minimized when e lies \perp to column space of A
- Thus we seek w such that $A^T e = 0 = A^T (y Aw)$.

• Thus
$$A^T y - A^T A w = 0$$
.

- Thus $A^T y = A^T A w$ (same as from MLE/least squares)!
- Thus $w = (A^T A)^{-1} A^T y$, as before.

 $\begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix} = A_W = \begin{pmatrix} x_1 \\ x_n \\ \vdots \\ y_n \end{pmatrix} W$

$$\mathbb{R}^{N}$$

ONTO

n space of A

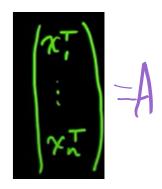
canachieve this

Using basis expansions instead of x_i

- $\{\mathbf{x}_j\} \rightarrow \{\Phi(x_j)\}$?
- Just define A with $\Phi^{T}(x_{j})$ instead of x_{j} because $\Phi(x)$ is fixed ahead of time, so its like someone just gave us different raw inputs x.

$$x = \Phi[x_1, x_2] = [1, x_1, x_2, x_1x_2, x_1^2, x_2^2] \in \mathbb{R}$$

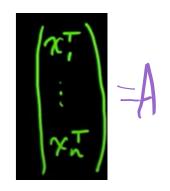
Polynomial expansion of order 2 (i.e. quadratic)



What can go wrong in linear regression

When can we invert $A^T A \in \mathbb{R}^{d \times d}$ (for *d* features)? $(A \in \mathbb{R}^{N \times d})$

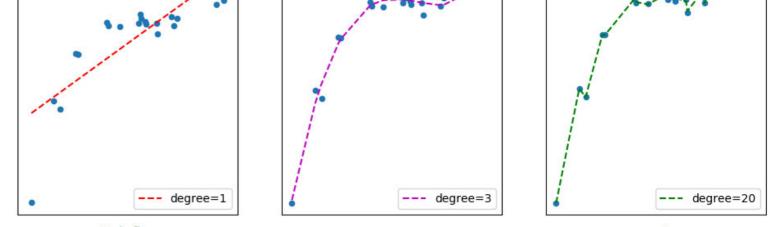
- Consider the decomposition we learned last class:
- $A^T A = \Sigma = Q D Q^T$ for diagonal D containing eigenvalues, and orthonormal Q.
- Then we had that $(A^T A)^{-1} = Q D^{-1} Q^T$. So if $A^T A$ has any zero eigenvalues we cannot invert it (there are ∞ equally good solutions w then).
- (Could use Moore-Penrose pseudo inverse.)
- This degeneracy occurs when the features in A, for the given data $\{x_j\}$ are linearly dependent.
- e.g., when d > n (then $rank(A^TA) < n$, but needs to be D, i.e. "full rank" to avoid zero eigenvalues).
- What about when d = n?



What can go wrong in linear regression

As we add higher and higher order polynomials, a few things happen:

- 1. # features, d gets bigger and bigger.
- 2. For $d \ge n$ can perfectly fit any data (*i.e.*, polynomials are a complete basis).
- 3. Even when we don't perfectly fit the training data, we are still in danger of overfitting (worse prediction on test set). Our goal is not to fit a line through the training data exactly, it is to do well on unseen test cases!



https://towardsdatascience.com/polynomial-regression-bbe8b9d97491

What can go wrong in linear regression

Two main categories of fixes:

- 1. Remove features until the problem is well behaved.
- 2. Leave the features as they are, but add constraints to the system to "tighten it up" (aka "regularization").
- 1) Is called "feature selection", e.g. "forward selection", "backward selection", etc.

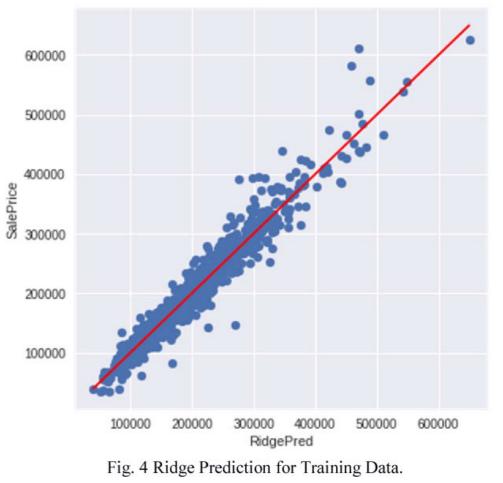
(Moore-Penrose inverse is not a general fix for ML models, only works for linear regression.)

A thought experiment

Consider:

- Use MLE on data, $D = \{x_i, y_i\}$ to get $\widehat{y}_i = p_{\theta}(y | \Phi(x))$ using linear regression.
- Assume an abundance of data (millions of data points), and only 100 parameters.
- Suppose get accuracy ∓\$1000 of sale price when applying to held out part of our data.
- Can we assume this model will get ∓\$1000 on any test set that may come in the future?

Actual vs. predicted sale price of house



Causation vs correlation

Breakingviews

Zillow's failed house flipping

Reuters

WSJ NOV. 2021 : "The company expects to record losses of more than \$500 million from homeflipping by the end of this year and is laying off a quarter of its staff."

Actual vs. predicted sale price of house

